



Spatial distribution and sampling plans for Grape vine plant canopy inhabiting *Scaphoideus titanus* (Hemiptera, Cicadellidae) nymphs

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3 Spatial distribution and sampling plans for Grape vine plant canopy
4 inhabiting *Scaphoideus titanus* (Hemiptera, Cicadellidae) nymphs

5

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25 **ABSTRACT** The paper deals with the study of the spatial distribution and the design
26 of sampling plans for estimating nymph densities of the Grape leafhopper
27 *Scaphoideus titanus* Ball in vine plant canopies. In a reference vineyard, leaf
28 samples were repeatedly taken according to a multi-stage stratified random sampling
29 procedure and data were subjected to an ANOVA. There were no significant
30 differences in density neither among the strata within the vineyard nor between the
31 two strata with basal and apical leaves. The significant differences between densities
32 on trunk and productive shoots lead instead to the adoption of two-stage (leaves,
33 plants) and three-stage (leaves, shoots, plants) sampling plans for trunk shoots and
34 productive shoots inhabiting individuals respectively. The mean crowding to mean
35 relationship used to analyze the nymphs spatial distribution revealed aggregated
36 distributions. In both the enumerative and the sequential enumerative sampling
37 plans, the number of leaves of trunk shoots, and of leaves and shoots of productive
38 shoots, was kept constant while the number of plants varied. In additional vineyards
39 data were collected and used to test the applicability of the distribution model and the
40 sampling plans. The tests confirmed the applicability i) of the mean crowding to mean
41 regression model on the plant and leaf stages for representing trunk shoot inhabiting
42 distributions, and on the plant, shoot and leaf stages for productive shoot inhabiting
43 nymphs, ii) of the enumerative sampling plan, iii) of the sequential enumerative
44 sampling plan. In general, sequential enumerative sampling was more cost-efficient
45 than enumerative sampling.

46

47 **KEY WORDS** enumerative sampling plan, sequential enumerative sampling plan,
48 mean crowding to mean regression, reliability levels, simulation

49

50 The North American leafhopper *Scaphoideus titanus* Ball (Hemiptera: Cicadellidae)
51 arrived in Europe in the 1950s to reach quickly the status of key pest in European
52 vineyards (Bonfils and Schwester 1960, Schwester et al. 1961, 1962a, 1962b). It
53 invaded Southwestern and Southern European grapevine growing areas from where
54 it spread to Southeastern European regions. Currently, the area of distribution
55 spreads from Portugal to Bulgaria and is still expanding (Chuche and Thiery 2014).
56 *Scaphoideus titanus* is the vector of the *Candidatus* Phytoplasma vitis, a
57 Phytoplasma of the Elm Yellows or 16Sr-V group, an economically important A2
58 quarantine pest for EPPO causing the Flavescence dorée (FD). To plan and carry
59 out control operations, the occurrence of FD and its vector is monitored throughout
60 invaded and prospective areas of distribution (EPPO/CABI 1996).

61 Beating tray and sticky trap techniques are satisfactory for monitoring *S. titanus*
62 nymph and adult occurrences in the vine plant canopy and useful for validating short
63 and long term phenology models (Rigamonti et al. 2011, 2013, 2014) but inadequate
64 for obtaining intensities and absolute densities. Intensity estimates for vine plant leaf
65 inhabiting life stages obtained through visual examination of leaves proved to be
66 useful for between plant distribution studies, sampling plan design and canopy
67 infestation model development (Lessio and Alma 2006, Maggi et al. 2013). To obtain
68 absolute density estimates for nymphs, the proportion of nymphs inhabiting the
69 vegetation of the vineyard floor and the number of leaves per plant has to be taken
70 into account (Trivellone et al. 2011, 2013). Intensity estimates have long been used
71 for analysing spatial distributions in heterogeneous sampling universes (Southwood
72 1978). They improve the insight into population processes in spatially structured
73 environments and contribute to further developing supervised vector management
74 programs (Pedigo and Buntin 1994).

75 Already in the 1950s, LeRoux and Reimer (1959) used stratified multi-stage
76 sampling procedures in apple orchards. They randomly selected subsamples at
77 different stages that they stratified for taking into account stage-specific spatial
78 heterogeneities. The procedures have been detailed by Cochran (1977) and recently
79 applied to spatial studies of vineyard insects by Ifoulis and Savopoulou-Soultani
80 (2006). LeRoux and Reimer (1959) showed how to analyze the variance of within-
81 and between plant components and opened the door to the development of spatial
82 distribution models applicable to highly structured sampling universes, in that the
83 variance or related statistics are often regressed on mean intensity (Taylor 1961,
84 Iwao 1968, Kuno 1976). Since these models take into account the dependency of the
85 variance on the mean density, they are generally valid on a wide range of densities
86 (Davis, 1994).

87 For the estimation of the parameters of these models, quite large sets of data are
88 required. Since, for a given scale, the parameters tend to be species-specific and
89 constant through time and space, researchers rationalize data collection by applying
90 repeated sampling procedures (Iwao and Kuno 1971, Hagihara 1976, Taylor and
91 Woiwood 1980, Perry and Taylor 1986, Sawyer 1989).

92 The purpose of this paper is to apply multi-stage stratified random sampling
93 procedures to analyze the spatial distribution patterns of *S. titanus* in a highly
94 structured grape vine plant canopy of a FD free reference vineyard. Knowledge on
95 the sources of spatial variations can be used to set the stage for adapting Kuno's
96 (1976) regression model and employing it for modeling the spatial distribution and
97 designing sampling plans. A successful application of the distribution model and the
98 sampling plans in additional vineyards is expected to improve the sampling strategy
99 and to put adaptive *S. titanus* management in FD free areas on a more solid ground
100 than available so far.

101

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Materials and Methods

103

104 **Study vineyards.** In a reference vineyard located at Camorino (46°09' N, 8°59' E,
105 220 m a.s.l.) in Southern Switzerland, the number of *S. titanus* individuals was
106 recorded repeatedly in 2009 to acquire data for analyzing the sources of variation,
107 distribution model development and sampling plan design.

108 Additional vineyards were sampled at Arbedo (46°13' N, 9°02' E, 240 m a.s.l.),
109 Bironico (46°07' N, 8°55' E, 490 m a.s.l.) and Contone (46°09' N, 8°56' E, 210 m
110 a.s.l.) to acquire data for testing the applicability of the distribution model and the
111 sampling plans

112 The plants were pruned to allow shoot development departing from 10 buds and
113 trained according to the Guyot double system. These shoots are referred to as
114 productive shoots that are different from the spontaneous unproductive trunk shoots.
115 Occasionally, the productive shoots were topped and hedged. Spontaneous
116 vegetation, occasionally mowed, covers the ground between the rows. Within the
117 rows, herbicides were used to control weeds. Synthetic pesticides were applied
118 against powdery and downy mildew but not against arthropods.

119 To facilitate the comparison with literature information, we simply refer to
120 'densities' rather than to 'relative densities' or 'intensities'. When referring to
121 subsampling, we refer to stages and levels in an interchangeable way.

122 **Sampling procedures.** LeRoux and Reimer's (1959) procedure takes into
123 account the spatial heterogeneity of the reference vineyard, the shoot type and the
124 within-shoot leaf position, and enhances the sampling efficiency through sub-
125 sampling the plant and the shoot. According to that procedure, to take into account
126 the possible within-vineyard heterogeneity, we stratified the reference vineyard to

127 obtain five plots of equal size, one at each side and one in the centre of the vineyard.
128 In each stratum, five plants were selected at random. Sub-sampling of the plant was
129 then carried out by randomly selecting three productive shoots and one trunk shoot.
130 To take into account a possible heterogeneity within productive and trunk shoots, all
131 shoots were divided into two equal strata, separating the leaves of the basal half from
132 those of the apical half. In each shoot stratum, three leaves were selected at random.
133 On each leaf, the number of *S. titanus* individuals was counted. The vineyard was
134 sampled weekly from 27 May 2009 to 30 July 2009. The samples taken on 27 May,
135 10 June, 17 June, 25 June and 3 July contained a sufficient number of nymphs for
136 conducting statistical analyses.

137 **Sources of variation.** To identify the sources of variation between sampling units
138 in the reference vineyard, analyses of variance (ANOVA) were carried out as
139 described by LeRoux and Reimer (1959). The between-plants variability is attributed
140 to the differences between the strata at the vineyard level and between the plants.
141 The within-plant variability is attributed to the differences between productive shoots
142 and the trunk shoots, and to the differences between the strata within the shoots. The
143 plant is considered as a random factor, as well as the productive shoots, while
144 vineyard strata, shoot type and shoot strata are considered as fixed. We followed
145 LeRoux and Reimer's (1959) recommendation of using untransformed counts.

146 **Spatial distribution model.** Iwao's (1968) model regresses Lloyd's (1967) mean
147 crowding index m^* on mean density m , i.e.

$$148 \quad m^* = \alpha + \beta m \quad [1]$$

149 where α and β have been defined as index of basic contagion and density-
150 contagiousness coefficient, respectively, and qualified as meaningful indexes for the
151 spatial characteristics of a population (Davis 1994). Thereby, the parameters are

152 related to the basic population unit and to the aggregation of population units; the
153 basic unit is a colony, an individual or individuals repelling each other if $\alpha > 0$, $\alpha = 0$ or
154 $\alpha < 0$, respectively, while the spatial distribution is aggregated, random or uniform if
155 $\beta > 1$, $\beta = 1$, $\beta < 1$, respectively. Theoretical and experimental information suggests that
156 the regression parameters are species specific and constant through time and space,
157 for a given scale and range of densities (Iwao and Kuno 1971, Hagihara 1976, Taylor
158 and Woiwood 1980, Perry and Taylor 1986, Sawyer 1989). According to Kuno
159 (1976), the model holds for habitats of any given size, for a whole area as well as for
160 sub-samples.

161 Kuno (1976) used model [1] to represent the distribution of insects in two- and
162 three stage random sampling programs. To facilitate the consultation of his paper, we
163 use whenever possible his symbols and terminology. Since the shoot types are the
164 most important source of variation (see the Results and Discussion section),
165 separate analyses for trunk shoot and productive shoot inhabiting populations are
166 carried out. The equation [1] then becomes

$$167 \quad m_j^* = \alpha + \beta_j m \quad \text{with } j=1,2 \text{ for trunk shoot distribution analyses} \quad [2]$$

$$168 \quad m_j^* = \alpha + \beta_j m \quad \text{with } j=1,2,3 \text{ for productive shoot distribution analyses} \quad [3]$$

169 In the case of trunk shoots, the model is applied on the level of the vineyard ($j=1$),
170 with plants as primary sampling units (PSU), and on the level of the plant ($j=2$), with
171 leaves as secondary sampling units (SSU) (eq. [2]). In the case of productive shoots,
172 the model is applied on the level of the vineyard ($j=1$), with plants as primary
173 sampling units (PSU), on the level of the plant ($j=2$), with shoots as secondary
174 sampling units (SSU) and on the level of the shoot ($j=3$) with leaves as the tertiary
175 sampling unit (TSU) (eq. [3]).

176 Empirical information suggests that individuals are the basic unit of the population.
177 Waters *et al.* (2014) showed that modifying both the linear function and the statistical
178 model to force the intercept or lower functional limit through the origin results in more
179 intuitive biological interpretation of parameters and better sampling economy. Hence,
180 the regression models were forced through the origin at all stages under
181 consideration ($\alpha=0$). The mean crowding values were calculated for between plants,
182 between shoots and, for productive shoots only, within shoot relationships. The
183 parameters β_1 and β_2 for trunk shoot inhabiting individuals, and β_1 , β_2 and β_3 for
184 productive shoots inhabiting individuals were obtained via least square regression
185 analyses.

186 **Reliability of density estimates and design of sampling plans.** The variance of
187 the density estimates for individuals living on trunk shoots and productive shoots can
188 be calculated with eqs. [2] and [6] in Kuno's (1976) paper. To introduce the
189 dependency of the variance to the mean, Kuno (1976) re-writes the variances in
190 terms of the relevant mean crowding m^* to mean regression models (eqs. [2, 3]). For
191 assessing the reliability of ecological population estimates, the ratio D of the standard
192 error d to the mean m is used (Karandinos 1976, Kuno 1976, Pedigo and Buntin
193 1994).

194 The number of sampled units (leaves, shoots, plants) is small relative to the total
195 number of units, hence there is no need to make finite population corrections
196 (Cochran 1977), and Kuno's (1976) eqs. [16] and [25] can be used to develop
197 enumerative sampling plans.

198 For attaining a predefined reliability level D_0 with a fixed number k_0 of leaves per
199 plant, the number l of plants to be sampled, expressed as a function of the mean m
200 per leaf of individuals on trunk shoot leaves, is

$$l = \frac{\frac{1}{k_0} \left\{ \frac{\alpha+1}{m} + \frac{\beta_1(\beta_2-1)}{\beta_2} \right\} + \frac{\beta_1-\beta_2}{\beta_2}}{D_0^2 + \frac{1}{L} \frac{\beta_1-\beta_2}{\beta_2}} \quad [4]$$

where α is the intercept, and β_1 and β_2 are the slopes of eq. [2], and in our study $k_0=6$. For attaining a predefined reliability level D_0 with a fixed number k_0 of leaves per shoot and a fixed number q_0 of shoots per plant, the number l of plants to be sampled, expressed as a function of the mean m per leaf of individuals on productive shoot leaves, is

$$l = \frac{\frac{1}{k_0} \left[\frac{\beta_1-\beta_3}{\beta_3} - \frac{\beta_1-\beta_2}{\beta_2} + \frac{1}{q_0} \left\{ \frac{\alpha+1}{m} + \frac{\beta_1(\beta_3-1)}{\beta_3} \right\} \right] + \frac{\beta_1-\beta_2}{\beta_2}}{D_0^2 + \frac{1}{L} \frac{\beta_1-\beta_2}{\beta_2}} \quad [5]$$

where α is the intercept, and β_1 , β_2 and β_3 are the slopes of eq. [3], and in our study $k_0=6$ and $q_0=3$.

The high total number L of plants in the vineyard under consideration reduces $1/L$ to zero and hence, the denominators to D_0^2 . The computations are further facilitated taking into account $\alpha=0$ as stated above.

The disregard of the finite population correction permits also the reference to Kuno's (1976) eqs. [27] and [34] to develop sequential enumerative sampling plans. According to this plan, the field work is terminated when the cumulative total count T_n , with n being the sample size, *i.e.* the number of leaves sampled, reaches a predetermined value.

The sequential estimation of the density $m=T_n/n$ per leaf of trunk shoot inhabiting individuals is based on a constant number k_0 of leaves per plant. T_n is calculated after

$$T_n = \frac{\alpha+1}{D_0^2 - \frac{1}{n} \left\{ \frac{\beta_1(\beta_2-1)}{\beta_2} + \frac{\beta_1-\beta_2}{\beta_2} k_0 \right\}} \quad [6]$$

221 where $n = l k_0$, l being the number of plants, and $k_0=6$ in our study. For the sequential
 222 estimation of the density $m=T_n/n$ of productive shoot inhabiting individuals, the
 223 constancy of leaf numbers per shoot (k_0) and of shoot numbers per plant (q_0) leads to
 224 calculate T_n as follows

$$225 \quad T_n = \frac{\alpha + 1}{D_0^2 - \frac{1}{n} \left\{ \frac{\beta_1(\beta_3 - 1)}{\beta_3} + \frac{\beta_1(\beta_2 - \beta_3)}{\beta_2\beta_3} q_0 + \frac{\beta_1 - \beta_2}{\beta_2} k_0 q_0 \right\}} \quad [7]$$

226 where $n = l q_0 k_0$, l being the number of plants, $q_0=3$ and $k_0=6$. In eq. [6] and eq. [7], the
 227 intercept α is set to 0.

228 **Testing the applicability of the distribution model and the sampling plans**

229 **Sampling procedures.** The sampling plans in eqs. [4, 5] were applied in the
 230 Arbedo, Bironico and Contone vineyards where 100, 90 and 20 plants were selected
 231 at random, respectively. The number of individuals was counted on six randomly
 232 selected trunk shoot leaves, if the total number of leaves per trunk equaled or
 233 exceeded 20. Trunks with a smaller number of leaves were disregarded and
 234 substituted by another trunk. Among productive shoots, three were selected at
 235 random on each plant, and nymphs were counted on six randomly selected leaves
 236 per shoot. The Arbedo vineyard was sampled on 24 May 2011, the Contone vineyard
 237 on 25 May 2011, and the Bironico vineyard on 12 June 2013 and on 1st July 2013.
 238 The visits yielded four data sets that are used here for testing the applicability of the
 239 spatial distribution model and the sampling plans.

240 **Applicability of the distribution model.** The sampling data obtained in the
 241 additional vineyards were used to estimate the mean crowding to mean regression
 242 statistics. Importantly, only a single sample was available in the Arbedo and Contone
 243 vineyards, whereas the Bironico vineyard was sampled twice. Since the four samples
 244 were treated separately, there were no replicates available for evaluating the slopes

245 at the plant stage. The samples were used to calculate via least square linear
246 regression analysis the parameters of eqs. [2, 3] with respective coefficient of
247 determination R^2 and standard errors. The comparison with the reference vineyard
248 statistics is made by evaluating the overlapping of confidence intervals at the
249 probability level $\alpha/2=0.05$ according to Payton et al. (2003).

250 **Applicability of the sampling plans.** To test the applicability of the enumerative
251 and sequential enumerative sampling plans, we calculated the precision levels on the
252 basis of the aforementioned four data sets, as detailed below. In the following the
253 index c was used to denote the precision levels (D_c) obtained.

254 To verify the applicability of the enumerative sampling plans, we computed
255 hypothesized densities based on unpublished monitoring data collected in previous
256 years in the additional vineyards by collaborators of the Agricultural Bureau of
257 Canton Ticino and of the Agroscope centre of excellence for agricultural research in
258 Switzerland. The resulting hypothesized densities were 0.36 and 0.26 individuals per
259 leaf on trunk shoots and on productive shoots, respectively. For these densities the
260 optimum sample size (OSS) for the $D_0=0.1$ and $D_0=0.3$ reliability levels, according to
261 eqs. [4, 5], were 77 and 9 plants for trunk shoots and 69 and 8 plants for productive
262 shoots respectively.

263 For each OSS, we carried out 1000 simulations with randomly selected plants in the
264 Arbedo, Bironico and Contone data sets, and estimated the densities. Based on
265 these densities and on the slopes β_l of the distribution model for the reference
266 vineyard, 1000 reliability values D_c were then computed, to finally yield average
267 values. The limited data set available for the Contone vineyard allowed the
268 estimation of means and reliability levels D_c for productive shoots at $D_0=0.3$ only. For
269 testing the enumerative sampling plan, the mean values D_c are compared with the
270 predefined $D_0=0.1$ and $D_0=0.3$ reliability levels.

271 To verify the applicability of the sequential enumerative sampling plans the four
272 data sets were used to simulate 1000 times the sequential sampling of nymphs on
273 trunk and productive shoots at the two levels of reliability. In each of the 1000
274 simulations, the number of plants required to meet or exceed the stop line for
275 estimates on trunk and productive shoots was obtained. As for the enumerative
276 sampling plan, based on the estimates of the density and on the slopes β_l of the
277 distribution model for the reference vineyard, 1000 reliability values D_c were then
278 computed, to finally yield average values. For testing the reliability of the sequential
279 sampling plan, the average D_c were compared with the predefined $D_0=0.1$ and $D_0=0.3$
280 reliability levels. The limited data set available from the Contone vineyard allowed the
281 estimation of means and reliability levels D_c for productive shoots at $D_0=0.3$ only.

282 For testing the efficiency of the sequential enumerative sampling plan in relation to
283 the enumerative sampling plan, the average number of plants examined for reaching
284 or exceeding the stop line was calculated and compared to the number used in
285 enumerative sampling plan.

286 **Statistical analysis.** All the statistical analysis were conducted using the IBM SPSS
287 Statistics 22 software (IBM Software Group 2014, Chicago, IL).

288

289

Results and Discussion

290

291 **Sources of variation in reference vineyard.** According to Table 1, there were no
292 significant differences between vineyard strata, indicating a homogeneous
293 environment between plants. However, there was a significant difference between
294 shoot types at the beginning of the growing season. This is explained by the behavior
295 of nymphs. Newly hatched nymphs mainly inhabit trunk shoot leaves, while older
296 nymphs disperse in the plant canopy, often after a passage through the vineyard

297 floor vegetation (Trivellone et al. 2011, 2013). Unexpectedly, there was no significant
298 difference within the shoot strata. Possibly, the topping of productive shoots induces
299 the growth of secondary shoots and creates a sampling environment in that the
300 stratification produced two subdivisions that differ in their distance from the trunk
301 rather than grouping leaves according to their age, indicating that shoots form a
302 homogeneous sampling environment.

303 Furthermore there were no significant differences between the grapevine plants.
304 This was unexpected because the within-plant variance is often smaller than the
305 between plant variance, allowing researchers to focus on plant samples and limit the
306 number of within-plant samples (Southwood 1978, Lessio and Alma 2006).
307 Nevertheless, the plants as primary units are retained in this and future works,
308 because the interactions between grape vines, FD and *S. titanus* presumably occur
309 on the plant level.

310 Based on the result of the ANOVA the stratification of the vineyard and the shoots
311 could be disregarded in regression model development and in the design of the
312 sampling plans. However, the difference between shoot types required a
313 differentiation between them and to work with two-stage sampling plans for trunk
314 shoot inhabiting individuals and with three-stage sampling plans for productive
315 shoots inhabiting individuals.

316 The mean densities per leaf reported in Table 1 are similar to the ones observed
317 by Lessio and Alma (2006) and by one of the authors of this paper (M. Jermini,
318 personal communication) in many untreated vineyards with comparable architectures
319 in Northwestern Italy and Southern Switzerland, respectively. Because of the
320 occurrence of similar densities, canopy structures and management procedures, the
321 Camorino reference vineyard is seen as representative for the many vineyards in an
322 extensive FD-free grape growing region and possibly beyond.

323 **Reference vineyard spatial distribution model.** Figures 1 and 2 depict the mean
324 crowding to mean relationship at the plant and leaf level for trunk shoot inhabiting *S.*
325 *titanus* and at the plant, shoot and leaf level for productive shoot inhabiting *S.*
326 *titanus*, respectively. By visual examination, Iwao's (1976) linear regression models,
327 with intercept α set to 0, appeared to describe satisfactorily the relationships. The
328 slopes β_1 , β_2 and β_3 (in Fig. 2C only) were greater than 1 indicating an aggregated
329 distribution of population individuals as previously observed by Lessio and Alma
330 (2006). The aggregation on trunk shoot leaves was not unexpected since nymphs
331 may first aggregate in a highly heterogeneous vineyard floor to feed on preferred
332 plants and move to nearby vine plants thereafter (Trivellone et al. 2011, 2013).

333 **Reference vineyard sampling plans.** Figures 3 and 4 represent the enumerative
334 sampling plan for *S. titanus* individuals inhabiting trunk shoots and productive shoots,
335 respectively. The sampling plan referring to plants requires the taking of six leaves
336 per plant for trunk shoots or six leaves on each of three shoots for productive shoots.
337 The number of plants (l) depended on the density and the level of precision set to
338 $D_0=0.1$, adequate for research purposes, or $D_0=0.3$, suitable for pest management
339 activities.

340 This analysis deals with data obtained in vineyards in a FD free area. Presumably
341 the densities occurring in FD invaded areas are much smaller, leading to excessive
342 sample sizes for reliable population density estimates. For the time being this
343 possible disadvantage has no practical implications because in FD invaded areas the
344 mandatory treatments are undertaken without reference to population density
345 estimates.

346 Figures 5 and 6 represent the stop lines for the sequential enumerative sampling
347 for estimating the trunk shoot and productive shoot inhabiting individuals,
348 respectively. The optimum sampling plan referring to plants is designed for the

349 $D_0=0.1$ and $D_0=0.3$ reliability levels and requires the examination of six leaves per
350 plant or six leaves on each of three shoots for trunk shoot or productive shoots
351 inhabiting individuals, respectively.

352 The users of the sequential enumerative sampling plan have to control a minimum
353 number of plants, corresponding to the point at which the stop line reaches a positive
354 value, and, if required, continue to examine randomly selected additional plants until
355 the cumulative number of nymphs reaches the stop line. The minimum number of
356 plants depends on the shoot type and the level of precision D_0 and is 31 (for $D_0=0.1$)
357 or 4 (for $D_0=0.3$) when sampling trunk shoot inhabiting individuals, and 48 (for
358 $D_0=0.1$) or 6 (for $D_0=0.3$) when sampling productive shoot inhabiting individuals. In
359 both cases, the sequential method appears to be more efficient than the enumerative
360 plan since it often leads to a reliable estimate with a smaller number of examined
361 leaves.

362 **Spatial distribution in additional vineyards.** Table 2 shows the slopes with
363 associated standard errors of the mean crowding to mean regressions at the different
364 stages for trunk shoot and productive shoot inhabiting nymphs in the Bironico,
365 Contone and Arbedo vineyards. The slopes representing Iwao's (1968)
366 contagiousness show an aggregated distribution and moreover, little variability
367 among the samples. The highest degree of aggregation occurred on the shoot stage
368 for productive shoots, whereas the lowest degree of aggregation existed between
369 leaves on trunk shoots. The degree of aggregation was intermediate on leaves of
370 productive shoots.

371 The comparison of the average slopes given in Table 2 with means and
372 confidence intervals of the reference data set reported in Table 3 shows that all
373 means reported in Table 2 fall into the confidence interval obtained from Table 3.
374 Moreover, the confidence intervals given in Table 2 overlap with the confidence

375 intervals for the reference data set reported in Table 3. Hence, the parameter
376 estimates obtained from the Arbedo, Bironico and Contone can be considered in
377 agreement with the ones obtained in the Camorino reference vineyard.

378 Though the slope for densities on trunk shoots cannot be evaluated at the
379 vineyard stage (Table 2), it can be concluded that the spatial distribution model is
380 appropriate for the Arbedo, Bironico and Contone vineyards.

381 **Applicability of the sampling plans.** Table 4 shows that the number of plants,
382 calculated by applying the enumerative sampling plan to a hypothesized density,
383 yielded density estimates with mean reliability levels (columns 8 and 9) close to ones
384 predefined in the sampling plan (columns 4 and 5). This result was unexpected since
385 the hypothesized density (columns 2 and 3) used for determining the OSS was
386 generally different from the mean density obtained through simulations (column 6
387 and 7). Apparently, the enumerative sampling plans yielded density estimates at
388 satisfactory reliability levels even when using inaccurate hypothesized densities to
389 define the OSS. This is due to generally similar and relatively high densities observed
390 in the Arbedo, Bironico and Contone vineyards, a density range within that the OSS
391 changes little.

392 Table 5 shows that the average reliability levels D_c obtained through 1000
393 simulations are in general close to the predefined reliability levels D_o , from which we
394 can conclude that the sequential enumerative sampling plan with the statistics
395 obtained from the Camorino reference data set yielded reliable estimates for the
396 Arbedo, Bironico and Contone vineyards.

397 Table 5 also shows that, at least for the trunk shoot inhabiting nymphs, the number
398 of plants required to obtain a reliable density estimate was often but not always lower
399 than the number of plants used in the enumerative sampling plan. Hence, the
400 sequential enumerative sampling plan was in general more efficient than the

401 enumerative sampling plan. Table 5 also shows that the sequential enumerative
402 sampling plan is less efficient if the hypothesized density is much higher than the
403 estimated density.

404 The methodology presented shows how sampling procedures in a heterogeneous
405 sampling universe can be structured and prepared for spatial distribution modeling.
406 The study of the sources for the spatial variation within the canopy led to the design
407 of a multi-stage stratified random sampling program that could be simplified further
408 into a two- and a three stage random sampling program that did not require
409 stratification any more.

410 Nevertheless, a brief review of the modern literature suggests that the approach
411 outlined by Pottinger and Le Roux (1971), Kuno (1976), Cochran (1977) and
412 Southwood (1978) three decades ago did not receive sufficient attention in spatial
413 analyses and sampling plan design. Without studying the sources of spatial variation
414 in detail, the authors of papers on regression models rely on assumptions to proceed
415 to model development.

416 The mean crowding to mean regression model used at different stages in a highly
417 structured vine plant canopy appears to be appropriate to describe the distribution of
418 nymphs in the reference and additional vineyards. We presume that the model is also
419 applicable to additional vineyards with similar canopy architecture, spontaneous
420 vegetation and surroundings, but we see a need to evaluate the applicability in
421 vineyards with different characteristics. The slopes of the mean crowding to mean
422 regressions indicated an aggregated distribution on the plant, shoot and leaf levels.
423 The results obtained here may provide some support to the claims that the
424 regression parameters are constant through time and space, for a given scale and
425 range of densities (Iwao and Kuno 1971, Hagihara 1976, Taylor and Woiwood 1980,
426 Perry and Taylor 1986, Sawyer 1989), and the model holds for habitats of any given

427 size, for a whole area as well as for sub-samples (Kuno 1976). Undoubtedly, more
428 data should be analyzed to confirm the validity of this claim.

429 Moreover, the results show that both the enumerative and the sequential
430 enumerative sampling plans can be applied to other vineyards with similar canopy
431 architecture, spontaneous vegetation and surroundings. To be able to apply
432 efficiently the enumerative sampling plan, the user should have some information on
433 possible densities in the vineyard to be sampled. This information doesn't have to be
434 accurate, however, and could be acquired in a preliminary sampling program or, as
435 done here, by approximating densities on the basis of previously conducted
436 monitoring programs. The results show that satisfactory levels of reliability are
437 reached even if the density estimate for determining the OSS is considerably
438 different from the actual density. This result may be of interest to pest managers
439 favoring enumerative over sequential enumerative sampling plans. As stated in the
440 case of the distribution model, the results obtained here should be critically evaluated
441 for other populations in different environments before drawing general conclusions.

442 The number of samples to be taken in the enumerative sampling program is higher
443 than the number reported by Lessio and Alma (2006). To some extent, the difference
444 may be due to different reliability levels for density estimates and the use of different
445 regression models that both can be justified (Davis 1994). Indeed, in contrast to this
446 work, Lessio and Alma (2006) relied on an application of Taylor's (1961) power law.
447 To a greater extent, however, the difference is due to the spatial variation under
448 consideration. Whereas Lessio and Alma (2006) neglected within-plant sources of
449 variability and focused on between-plant variation, the present findings show that the
450 within plant variability is more important than the between-plant variability and should
451 consequently be taken into account.

452 The sequential enumerative sampling plan can be applied to other vineyards with
453 a similar canopy architecture, spontaneous vegetation and surroundings. As
454 expected the sequential enumerative sampling plan is in general, but not always,
455 more cost efficient than the enumerative sampling plan. In sequential enumerative
456 density estimation, the applicability of the sampling plan doesn't require previous
457 knowledge on infestation levels and the effort to estimate population density
458 promises to be smaller than in the case of enumerative sampling. If, however, the
459 accumulated number of samples exceeds the one found in the enumerative sampling
460 plan, the user is advised to terminate the sampling activities and rely on the reliability
461 of the enumerative sampling plan.

462 The successful application and the applicability of the sampling plans in additional
463 vineyards put adaptive *S. titanus* management on more solid grounds than available
464 so far. So far, the *S. titanus* adaptive management project emphasized the within
465 vineyard population dynamics and the development of respective adaptive
466 management strategies. For this purpose the enumerative and sequential
467 enumerative sampling plans presented here are particularly appropriate and a
468 reliability level of $D_0=0.1$ was considered adequate. In the context of regional
469 population studies and disease transmission, however, the pest manager may be
470 interested in regional sampling and management operations. For this purpose, we
471 tentatively proposed a $D_0=0.3$ reliability level. As shown already by Sawyer and
472 Haynes (1985), regional population estimations require the extension of the spatial
473 model and the possible consideration of additional strata at the landscape level.

474 Further improvements of the sampling strategy for adaptive *S. titanus*
475 management can be made when re-orienting the work from research to adaptive
476 management under specific agricultural conditions. For example, improvements are
477 possible when focusing on initial densities on trunk shoots to decide on the necessity

478 of including the vineyard in a monitoring program. If a threshold density for this
479 decision can be given and the nymph density is at moderate levels, cost efficient
480 presence-absence sampling plans with respect to a threshold can be developed and
481 implemented (Wilson *et al.*, 1983; Bianchi *et al.*, 1989, Knapp *et al.*, 2006). Further
482 improvements are possible when focusing on specific developmental stages whose
483 occurrence may trigger a chemical control operation (e.g. in Switzerland the
484 appearance of third instar nymphs is used to determine the timing of the first
485 mandatory treatment). Guided by the results of continuously adapted phenology
486 simulations, the pest manager should be able to reduce the frequency of sampling
487 operations (Jermini *et al.*, 2013; Prevostini *et al.*, 2013).

488 The model describes the intensity per leaf of nymphs only. The analysis of the
489 spatial distribution of bark inhabiting eggs and mobile adults requires a different
490 sampling technique and a revision of the model structure. To obtain absolute density
491 estimates for all life stages from intensity estimates, the proportion of vineyard floor
492 inhabiting nymphs, which depends from the botanical composition, and the number
493 of leaves per plant has to be taken into account (Trivellone *et al.*, 2013).

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498 Bureau of the Canton Ticino kindly made available monitoring data from the Arbedo,
499 Bironico and Contone vineyards.

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617 **Table 1. Analysis of variance for visual *S. titanus* counts made on five sampling**
 618 **dates in a reference vineyard at Camorino (Switzerland) (DF = degrees of freedom, MS**
 619 **= observed mean squares, F = F-value; vineyard strata, shoot type and shoot strata**
 620 **are considered as fixed factors, while plants are considered as random factors; **:**
 621 **significant at $P < 0.01$)**

Sampling date	May 27	June 10	June 17	June 25	July 3						
(mean density)	(0.954)	(0.790)	(0.536)	(0.280)	(0.170)						
Source of variability	DF	MS	F	MS	F	MS	F	MS	F	MS	F
<i>Between plants</i>											
Vineyard strata	4	18.48	1.44	14.79	1.92	1.72	0.94	0.65	0.77	1.02	1.21
Plants	20	12.78	1.04	7.70	0.89	1.83	0.70	0.84	1.32	0.84	1.41
<i>Within plants</i>											
Shoot type	1	82.23	6.69**	178.20	20.64**	15.74	6.03**	0.04	0.06	1.62	2.07
Shoot strata	1	23.48	1.91	4.61	0.53	3.13	1.20	0.13	0.21	1.46	2.44
Residual	73	12.30		8.63		2.61		0.64		0.60	

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626 **Table 2. Mean crowding to mean regression statistics for *S. titanus* nymphs**
 627 **inhabiting trunk and productive shoots in the in the Arbedo, Bironico and Contone**
 628 **vineyards (for testing overlapping distributions, the confidence interval at the $\alpha/2=0.05$**
 629 **probability level is calculated after Payton et al. (2003)).**

Sample	Slope and confidence interval for trunk shoot inhabiting individuals		Slope and confidence interval for productive shoot inhabiting individuals	
	Stage 1	Stage 2	Stage 1	
	Between leaf relationship	Between shoot relationship	Between leaf relationship	
Arbedo vineyard	1.35 ± 0.1153	2.75 ± 0.4453	2.05 ± 0.1797	
Bironico vineyard, 12 June 2013	1.49 ± 0.1319	2.31 ± 0.3133	1.79 ± 0.1471	
Bironico vineyard, 1 st July 2013	1.36 ± 0.2144	2.34 ± 0.3958	1.95 ± 0.19661	
Contone vineyard	1.28 ± 0.0208	2.42 ± 0.7099	1.67 ± 0.2988	

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634 **Table 3. Mean crowding to mean regression statistics for *S. titanus* nymphs**
 635 **inhabiting trunk and productive shoots in the Camorino reference vineyard (for testing**
 636 **overlapping distributions, the confidence interval at the $\alpha/2=0.05$ probability level is**
 637 **calculated after Payton et al. (2003)).**

Stage	Slope	Standard error (SE)	R ²	n	Confidence interval
Trunk shoots					
Plant (β_1)	1.6784	0.1735	0.9045	5	0.3702
Leaf (β_2)	1.3603	0.7895	0.6344	79	1.3040
Productive shoots					
Plant (β_1)	2.8444	0.167	0.9563	5	0.3564
Shoot (β_2)	2.1325	0.960	0.6766	123	1.5777
Leaf (β_3)	1.8686	0.490	0.7398	296	0.8012

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642 **Table 4. The applicability of the enumerative sampling plan: the hypothesized mean**
 643 **densities with the required number of plants to be sampled, and means and reliability**
 644 **levels (mean values) obtained from 1000 simulated sampling procedures vineyards**
 645 **(nc = sample size too small for the calculations; the predefined reliability level D_0 is**
 646 **reported in parentheses).**

Vineyard	Hypothesized density on		Number of plants selected for		Estimated mean (1000 simulations) on		Reliability D_c (1000 simulations), mean values, on	
	trunk shoots	productive shoots	trunk shoot estimates	productive shoot estimates	trunk shoots	productive shoots	trunk shoots	productive shoots
Arbedo	0.36	0.26	77 (0.1)	69 (0.1)	0.96	0.22	0.09	0.11
			9 (0.3)	8 (0.3)	0.96	0.22	0.25	0.32
Bironico 12 June 2013	0.36	0.26	77 (0.1)	69 (0.1)	0.83	0.26	0.10	0.12
			9 (0.3)	8 (0.3)	0.83	0.26	0.30	0.34
Bironico 1 st July 2013 ^t	0.36	0.26	77 (0.1)	69 (0.1)	0.17	0.18	0.16	0.10
			9 (0.3)	8 (0.3)	0.17	0.18	0.43	0.28
Contone	0.36	0.26	nc (0.1)	nc (0.1)	nc	nc	nc	nc
			9 (0.3)	8 (0.3)	nc	0.18	nc	0.33

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650 **Table 5. The applicability of the sequential enumerative sampling plan: the most**
 651 **frequent number of plants required to be sampled for reaching or exceeding the stop**
 652 **line and the reliability levels (mean values) obtained from 1000 simulated sampling**
 653 **procedures (nc = sample size too small for the calculations).**

Vineyard	Shoot type	Most frequent number of plants required to meet the stop line	Calculated reliability (D_c), mean values
Predefined reliability $D_0 = 0.1$			
Arbedo	Trunk shoots	67	0.096
	Productive shoots	73	0.102
Bironico 12 June 2013	Trunk shoots	69	0.105
	Productive shoots	70	0.117
Bironico 1 st July 2013	Trunk shoots	149	0.113
	Productive shoots	78	0.087
Contone	Trunk shoots	nc	nc
	Productive shoots	nc	nc
Predefined reliability $D_0 = 0.3$			
Arbedo	Trunk shoots	8	0.246
	Productive shoots	9	0.284
Bironico 12 June 2013	Trunk shoots	8	0.296
	Productive shoots	9	0.308
Bironico 1 st July 2013	Trunk shoots	18	0.305
	Productive shoots	10	0.248
Contone	Trunk shoots	nc	nc
	Productive shoots	11	0.316

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657 **Fig. 1.** Mean crowding (m^*) to mean (m) regression statistics (density-
658 contagiousness coefficient β with SE = standard error, and coefficient of
659 determination R^2) for *S. titanus* nymphs inhabiting trunk shoots of grapevine plants in
660 the Camorino reference vineyard (A and B refer to the plant and leaf level,
661 respectively).

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664 **Fig. 2.** Mean crowding (m^*) to mean (m) regression statistics (density-
665 contagiousness coefficient β with SE = standard error, and coefficient of
666 determination R^2) for *S. titanus* nymphs inhabiting productive shoots of grapevine
667 plants in the Camorino reference vineyard (A, B and C refer to the plant, shoot and
668 leaf level, respectively).

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671 **Fig. 3.** Enumerative sampling plan with optimal sample size for *S. titanus* nymphs
672 inhabiting trunk shoot leaves. The solid and dotted lines represent the number of
673 plants to be sampled in relation to *S. titanus* density, at precision levels $D_0=0.1$ and
674 $D_0=0.3$ respectively.

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677 **Fig. 4.** Enumerative sampling plan with optimal sample size for *S. titanus* nymphs
678 inhabiting productive shoot leaves. The solid and dotted lines represent the number
679 of plants to be sampled in relation to *S. titanus* density, at precision levels $D_0=0.1$ and
680 $D_0=0.3$ respectively.

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683 **Fig. 5.** Sequential enumerative sampling plan for *S. titanus* nymphs inhabiting

684 trunk shoots (A: stop line for $D_0=0.1$; B: stop line for $D_0=0.3$).

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687 **Fig. 6.** Sequential enumerative sampling plan for *S. titanus* nymphs inhabiting

688 productive shoots (A: stop line for $D_0=0.1$; B: stop line for $D_0=0.3$).

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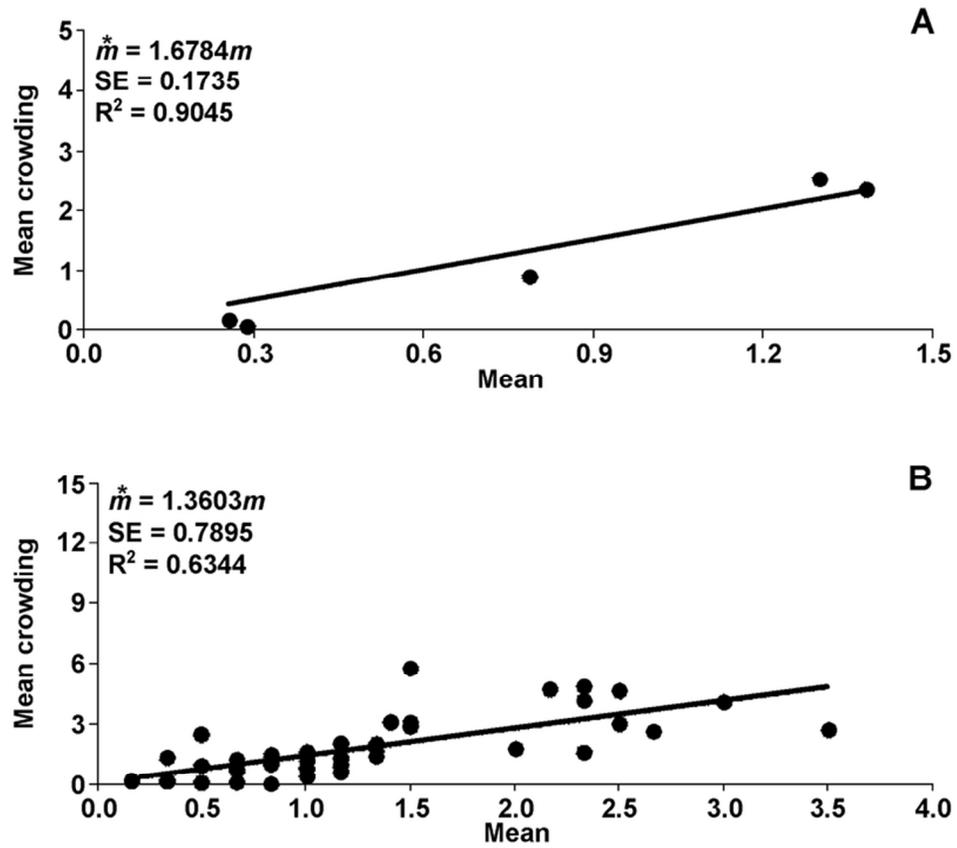


Fig. 1. Mean crowding (\hat{m}^*) to mean (m) regression statistics (density-contagiousness coefficient $\beta\beta$ with SE = standard error, and coefficient of determination R^2) for *S. titanus* nymphs inhabiting trunk shoots of grapevine plants in the Camorino reference vineyard (A and B refer to the plant and leaf level, respectively). 71x64mm (300 x 300 DPI)

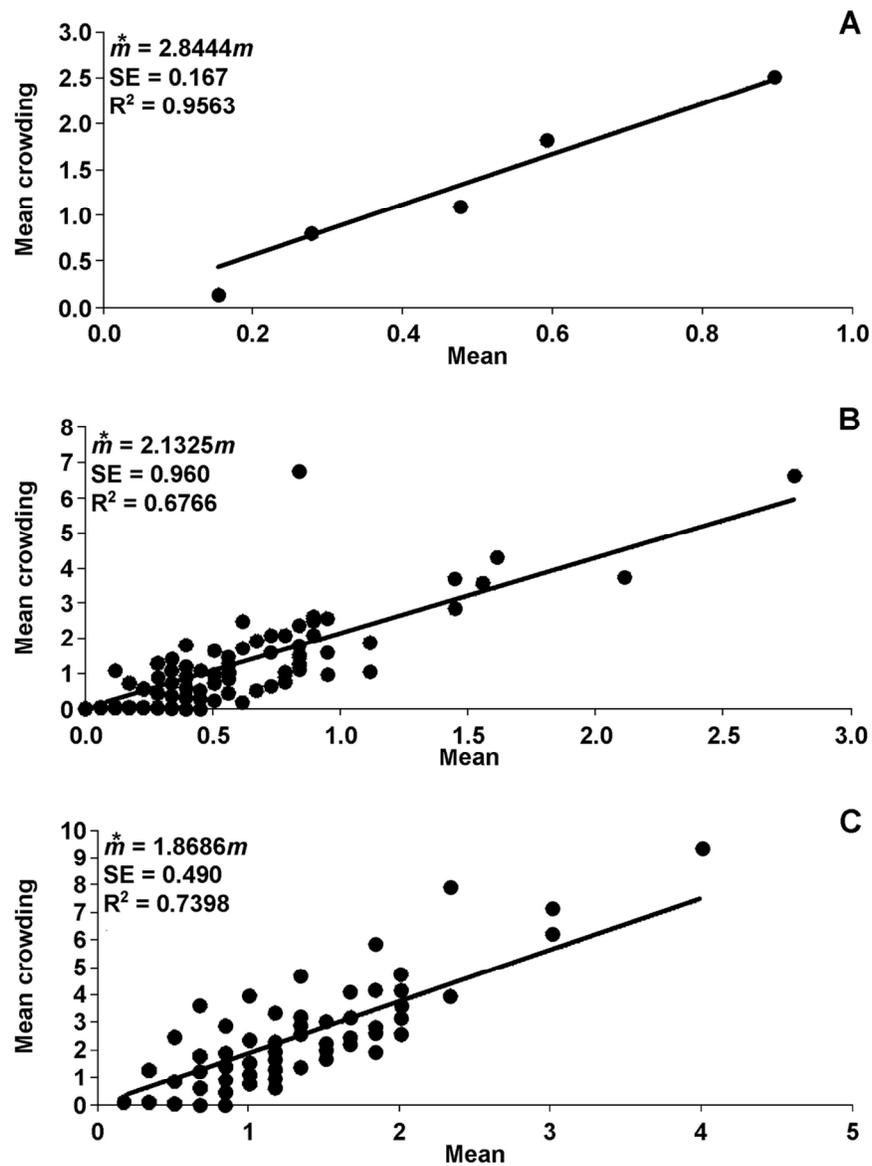


Fig. 2. Mean crowding (m^*) to mean (m) regression statistics (density-contagiousness coefficient β with SE = standard error, and coefficient of determination R^2) for *S. titanus* nymphs inhabiting productive shoots of grapevine plants in the Camorino reference vineyard (A, B and C refer to the plant, shoot and leaf level, respectively).

105x138mm (300 x 300 DPI)

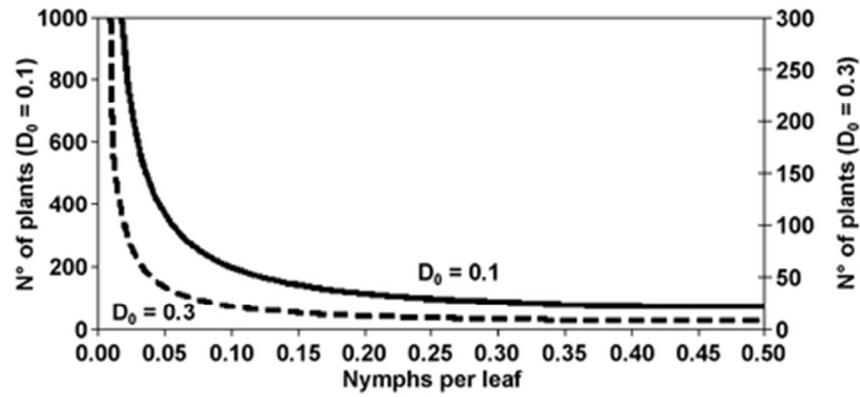


Fig. 3. Enumerative sampling plan with optimal sample size for *S. titanus* nymphs inhabiting trunk shoot leaves. The solid and dotted lines represent the number of plants to be sampled in relation to *S. titanus* density, at precision levels $D_0=0.1$ and $D_0=0.3$ respectively.
36x17mm (300 x 300 DPI)

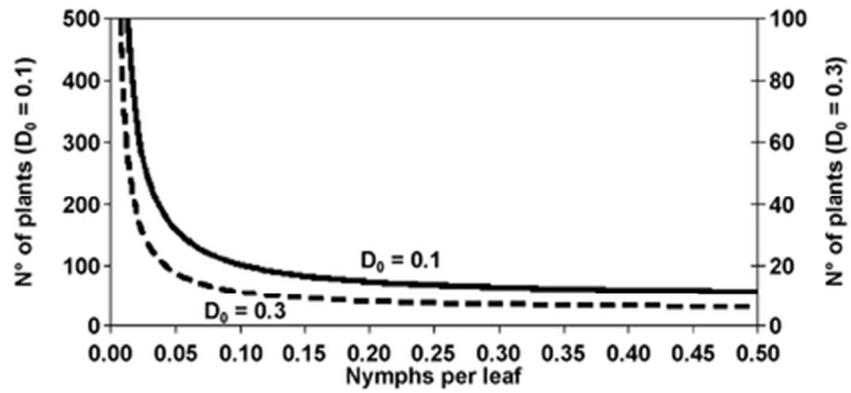


Fig. 4. Enumerative sampling plan with optimal sample size for *S. titanus* nymphs inhabiting productive shoot leaves. The solid and dotted lines represent the number of plants to be sampled in relation to *S. titanus* density, at precision levels $D_0=0.1$ and $D_0=0.3$ respectively.
36x16mm (300 x 300 DPI)

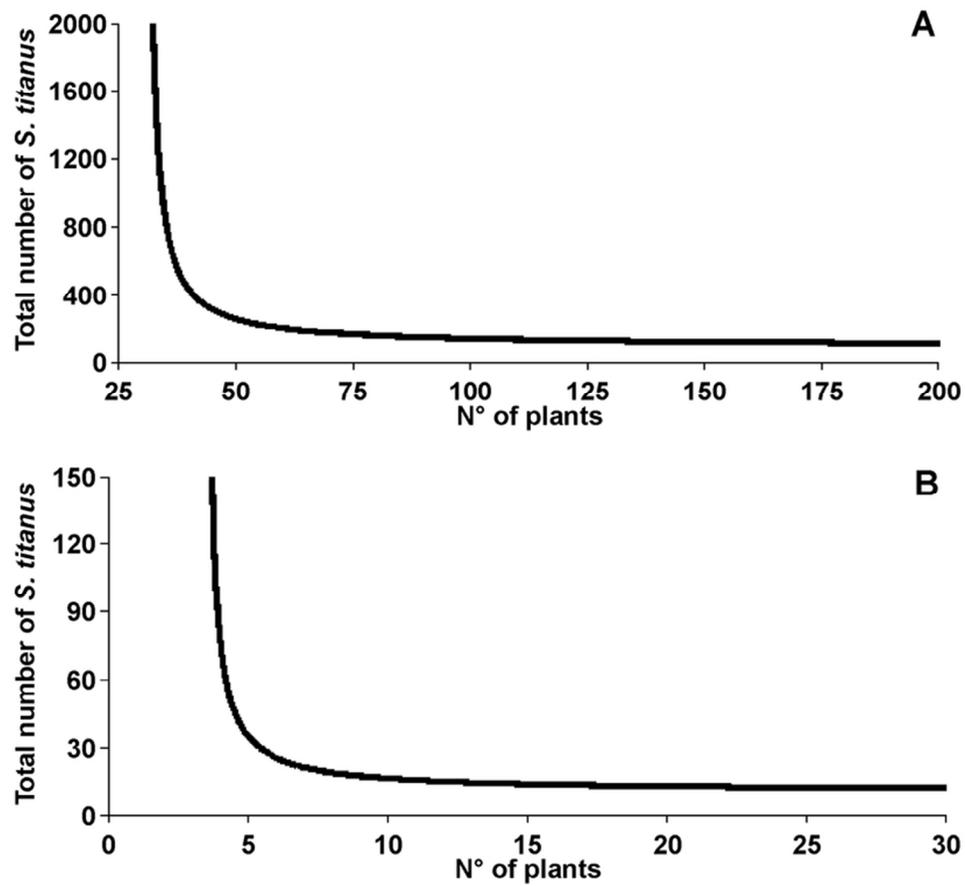


Fig. 5. Sequential enumerative sampling plan for *S. titanus* nymphs inhabiting trunk shoots (A: stop line for $D_0=0.1$; B: stop line for $D_0=0.3$).
72x65mm (300 x 300 DPI)

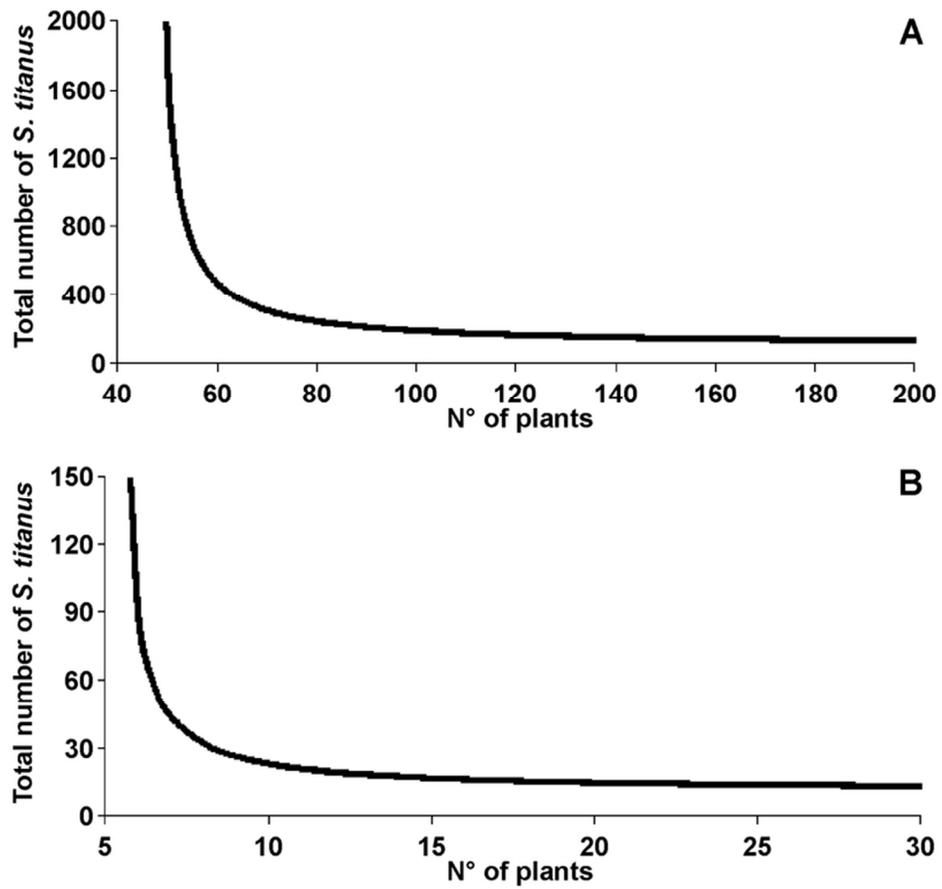


Fig. 6. Sequential enumerative sampling plan for *S. titanus* nymphs inhabiting productive shoots (A: stop line for $D_0=0.1$; B: stop line for $D_0=0.3$).
73x67mm (300 x 300 DPI)