# Weak measurements and the joint estimation of phase and phase diffusion 

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#### Abstract

Weak measurements offer the possibility of tuning the information acquired on a system, hence the imposed disturbance. This suggests that it could be a useful tool for multiparameter estimation, when two parameters cannot be measured simultaneously at the quantum limit. Here we discuss their use for phase estimation in the presence of phase diffusion in the context of polarimetry, a scenario that is conveniently cast in terms of a two-level quantum system in many relevant cases.


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## I. INTRODUCTION

Precision measurements often encounter intrinsic limitations imposed by quantum mechanics. Understanding these limitations and designing strategies for achieving the ultimate precision by means of quantum resources is the objective of quantum metrology [1]. The framework for single-parameter estimation is well established, both for Hamiltonian parameters, phases in particular [2-5], and for relevant cases of dissipative parameters [6,7].

Physical processes, though, show both unitary and dissipative dynamics. A possible approach is to treat dissipation as a stationary process, which can be characterized with arbitrary precision, and to consider phase estimation through the dissipative environment [8-15]. However, in the presence of nonstationary processes, a more satisfactory approach consists in the joint estimation of the parameters linked to the unitary part and the evolution as well as to the dissipation. This requires tackling the problem with the formalism for multiparameter estimation [16]. Concerning the case above, when a first parameter is ascribed to a unitary and a second to the dissipation, this has been applied to the cases of phase and loss [17] and phase with phase diffusion [18,19].

This latter case exemplifies the subtleties of multiparameter estimation. It has been recognized that there is no fundamental impediment to achieving the best possible precision for the phase shift $\phi$ and the magnitude of the phase diffusion $\delta$ simultaneously, and specific instances have been exemplified [18]. However, in practical cases, such as with coherent states, a tradeoff appears, which one cannot dispose of by using standard quantum resources [19], such as $N 00 N$ states or Holland-Burnett states [2]. This limitation appears to be intimately related to the geometric representation of these states, and only by carefully tailoring the quantum state can one circumvent these limitations in particular limits [18].

In this paper, we address this issue in polarimetry, in which the description is effectively carried out in terms of two-level systems, considering weak-measurement strategies. Their interest is in the possibility of repeated measurement of noncommuting observables [21,22]. There has been some debate about the scope of weak measurements in metrology, and where they could offer some practical advantages
[23-28]. Here we analyze how well weak measurements can be employed for multiparameter estimation, and we show under which conditions they saturate the tradeoff for the estimation of $\phi$ and $\delta$ with the best possible precision.

The paper is structured as follows: In Sec. II we present the theory behind multiparameter estimation of phase and phase diffusion, and its relationship with weak measurements. In Sec. III we discuss an experimental realization of a weak measurement device based on an interferometric scheme, discussing its performances in the estimation of the two parameters of interest. We conclude the paper in Sec. IV with remarks and possible outlooks.

## II. THEORY

Our ideal estimation experiment is illustrated in Fig. 1. Polarized light passes through a sample, which imparts a birefringent phase; this is distributed around an average value $\phi$, with a distribution whose variance is $\delta$ (which will be referred to as phase diffusion for brevity). The output density matrix, generated by the action of the channel on the optimal probe $|\psi\rangle=\frac{1}{\sqrt{2}}\left(\left|y_{+}\right\rangle+\left|y_{-}\right\rangle\right)$, can be written in the basis $\left\{\left|y_{ \pm}\right\rangle\right\}$on which the phase shift occurs as $[16,19]$

$$
\varrho=\frac{1}{2}\left(\begin{array}{cc}
1 & e^{-1 \phi-\delta^{2}}  \tag{1}\\
e^{1 \phi-\delta^{2}} & 1
\end{array}\right) .
$$

The performance of an unbiased estimator for the two independent parameters $\phi$ and $\delta$ is quantified by the covariance matrix

$$
\boldsymbol{\Sigma}=\left(\begin{array}{cc}
\operatorname{var}(\phi) & \operatorname{cov}(\phi, \delta)  \tag{2}\\
\operatorname{cov}(\phi, \delta) & \operatorname{var}(\delta)
\end{array}\right)
$$

The covariance matrix, obtained via a number $M$ of repetitions of the experiment, is lower bounded as predicted by the classical and quantum Cramér-Rao bounds as follows:

$$
\begin{equation*}
M \Sigma \geqslant \mathbf{F}^{-1} \geqslant \mathbf{H}^{-1} \tag{3}
\end{equation*}
$$

where the matrices $\mathbf{F}$ and $\mathbf{H}$ are called, respectively, Fisher information and quantum Fisher information (QFI) matrices [16] (we report in Appendix the general formulas for the two matrices). While the matrix $\mathbf{F}$ depends on the specific


FIG. 1. (Color online) Schemes for multiparameter metrology. A probe is prepared in the polarization state $|\psi\rangle=\frac{1}{\sqrt{2}}\left(\left|y_{+}\right\rangle+\left|y_{-}\right\rangle\right)$and sent in a sample imparting a random phase shift, whose distribution has a mean value $\phi$ and a characteristic width $\delta$. Two possible schemes are considered: a four-outcome POVM $\Pi_{x, z}$, composed of projectors along the eigenvectors of $\sigma_{x}$ and $\sigma_{z}$, and known to saturate the bound (4) as shown in [19]; and a weak measurement $M$ along the $\sigma_{z}$ axis, followed by a projective measurement $\Pi_{x}$ along $\sigma_{x}$.
measurement (POVM) performed, the QFI matrix $\mathbf{H}$ is a property of the output state $\varrho$ only. In this particular instance, $\mathbf{H}$ is diagonal, hence we can find separate bounds for each variance independently: $\operatorname{var}(\phi) H_{\phi \phi} \geqslant M$ and $\operatorname{var}(\delta) H_{\delta \delta} \geqslant M$ [18,19], where

$$
\begin{aligned}
H_{\phi \phi} & =e^{-2 \delta^{2}}, \\
H_{\delta \delta} & =\frac{4 \delta^{2}}{e^{2 \delta^{2}}+1}
\end{aligned}
$$

are the diagonal elements of $\mathbf{H}$ associated with $\phi$ and $\delta$, i.e., the QFIs associated with each parameter [19]. Although $\mathbf{H}$ is diagonal, so that the parameters are in principle statistically independent, these cannot be measured simultaneously at the quantum limit using a single two-level system, even when employing POVMs other than projectors and ensuring that no correlations between the two estimators arise [19]. This descends from the general result that in the multiparameter case, the quantum Cramér-Rao bound is not tight, and it may not be saturated [16]. In fact, with $\mathbf{F}$ denoting the Fisher information matrix associated with any possible measurement operator on a two-level system, the following bound holds:

$$
\begin{equation*}
\frac{F_{\phi \phi}}{H_{\phi \phi}}+\frac{F_{\delta \delta}}{H_{\delta \delta}} \leqslant 1 \tag{4}
\end{equation*}
$$

which, in turn, implies

$$
\begin{equation*}
\frac{F_{\phi \phi}^{\prime}}{H_{\phi \phi}}+\frac{F_{\delta \delta}^{\prime}}{H_{\delta \delta}} \leqslant 1 \tag{5}
\end{equation*}
$$

where we have introduced the effective Fisher information $F_{\phi \phi}^{\prime}=1 /\left(\mathbf{F}^{-1}\right)_{\phi \phi}=F_{\phi \phi}-F_{\delta \phi}^{2} / F_{\delta \delta}$ and $F_{\delta \delta}^{\prime}=1 /\left(\mathbf{F}^{-1}\right)_{\delta \delta}=$ $F_{\delta \delta}-F_{\delta \phi}^{2} / F_{\phi \phi}$. As a consequence of Eq. (3), these quantities are in fact the ones bounding the variances for each estimator, given a specific measurement strategy, i.e., $\operatorname{var}(\phi) F_{\phi, \phi}^{\prime} \geqslant 1 / M$ and $\operatorname{var}(\delta) F_{\delta, \delta}^{\prime} \geqslant 1 / M$ [16]. By using the effective Fisher information, then, one explicitly takes into account the correlations between the parameters, thus collecting information about one parameter limits the information accessible on the other. Since in principle the right-hand side of Eqs. (4) and (5) can be as high as 2, this bound establishes that there exists no better strategy than using a fraction of the experimental
runs to estimate the phase, and the rest for the phase diffusion, with a weighting giving the relative importance of the two parameters. A measurement saturating (4) is described by a rank-1 POVM, measuring along the eigenvectors of $\sigma_{x}$ and $\sigma_{z}$. In this measurement strategy, one can tune the relative weight of each parameter by tuning the probability of performing the $\sigma_{x}$ or the $\sigma_{z}$ measurement. We observe that this bound is related to a similar relation established in [20] for state estimation.

Weak measurements offer a different strategy for addressing this joint estimation problem. We can imagine a scheme in which $\sigma_{z}$ is first measured weakly, and then $\sigma_{x}$ is measured with a standard projection. The weak measurement provides on average an unbiased estimator [29], and although it always introduces a disturbance, this can be taken into account in the data analysis by means of some form of calibration, at the cost of a reduction in the attainable precision. We quantify the quality of this measurement by evaluating its Fisher information. To do so, we first introduce weak measurement operators in the form

$$
\begin{equation*}
M_{ \pm}=\frac{1}{\sqrt{2}}\left[\cos (\theta / 2) \sigma_{0} \pm \sin (\theta / 2) \sigma_{z}\right] \tag{6}
\end{equation*}
$$

where $\sigma_{0}$ is the $2 \times 2$ identity matrix. The angle $\theta$ establishes the amount of information extracted in a single run, and it ranges from 0 , corresponding to no measurement and no disturbance, to $\pi / 2$, corresponding to full projective measurement. These operators allow us to find the final detection probabilities using the modified Born's rule $p( \pm, \pm)=\frac{1}{2} \operatorname{Tr}\left[M_{ \pm} \varrho M_{ \pm}\left(\sigma_{0} \pm\right.\right.$ $\left.\left.\sigma_{y}\right)\right]$, and similarly in the other cases [29,30]. Our goal here is to relate the Fisher information $\mathbf{F}$ to the invasiveness parameters $\theta$ of the weak measurement; the Fisher information can be easily computed from the conditional probabilities $\{p(+,+), p(+,-), p(-,+), p(-,-)\}$ by using the formula

$$
\begin{gather*}
F_{\alpha, \beta}=\sum_{\{j, k\}= \pm} p(j, k)\left[\partial_{\alpha} \log p(j, k)\right]\left[\partial_{\beta} \log p(j, k)\right] \\
\text { with } \alpha=\{\phi, \delta\} \text { and } \beta=\{\phi, \delta\} \tag{7}
\end{gather*}
$$

In particular, we obtain

$$
\begin{gather*}
F_{\phi \phi}=\frac{e^{-2 \delta^{2}}}{2}\left(\frac{\cos ^{2}(\theta-\phi)}{1-e^{-2 \delta^{2} \sin ^{2}(\theta-\phi)}}\right. \\
\left.+\frac{\cos ^{2}(\theta+\phi)}{1-e^{-2 \delta^{2} \sin ^{2}(\theta+\phi)}}\right),  \tag{8}\\
F_{\delta \delta}= \\
2 \delta^{2} e^{-2 \delta^{2}}\left(\frac{\sin ^{2}(\theta-\phi)}{1-e^{-2 \delta^{2} \sin ^{2}(\theta-\phi)}}\right.  \tag{9}\\
+\frac{\sin ^{2}(\theta+\phi)}{\left.1-e^{-2 \delta^{2} \sin ^{2}(\theta+\phi)}\right),} \\
F_{\phi \delta}=F_{\delta \phi}=  \tag{10}\\
\frac{\delta e^{-2 \delta^{2}}}{2}\left(\frac{\sin (2 \theta-2 \phi)}{1-e^{-2 \delta^{2}} \sin ^{2}(\theta-\phi)}\right. \\
\\
-\frac{\sin (2 \theta+2 \phi)}{\left.1-e^{-2 \delta^{2} \sin ^{2}(\theta+\phi)}\right)} .
\end{gather*}
$$

Direct substitution verifies that our weak measurement scheme always saturates, but it cannot beat the bound (4). This fact can be understood by observing that the overall
measurement operator, of the form $M_{+}\left(\sigma_{0}+\sigma_{x}\right) M_{+}$, corresponds to a projector in the direction $(\sin (\theta), 0, \cos (\theta))$ of the Bloch sphere, and similar expressions hold in the other cases, originating a symmetry around the $y$ axis; thus, weak measurements are bound by the expression (4), and indeed saturate it (see the supplemental information of Ref. [19]). This bound, however, does not give prescriptions as to the behavior of two-parameter estimation in terms of the strength of the weak measurement, a question that we address in the following.

The saturation of Eq. (4) does not ensure that the scheme can actually provide two estimators for $\phi$ and $\delta$, both achieving the optimal bound for the effective Fisher information (5). This is because there might exist correlations between these two parameters introduced by the off-diagonal term $F_{\delta \phi}$. However, one can then directly check that for specific values of the phase, $\phi=k \pi / 2$ (with $k \in \mathbb{Z}$ ), these correlations are always equal to zero, and then the bound (5) is saturated as well for all values of the other parameter to be estimated $\delta$ and for all values of the measurement strength parameter $\theta$. By noticing that a phase rotation on the input state is completely equivalent to a rotation of the directions of both the weak measurement


FIG. 2. (Color online) Effective Fisher information as a function of the measurement strength $\theta$, rescaled to the corresponding values of the QFI for $\phi=0$ and for different values of $\delta$ (top panel, $\delta=0.1$; bottom panel, $\delta=1$ ). In each plot, the solid blue line corresponds to $F_{\phi \phi}^{\prime} / H_{\phi \phi}$ while the red dashed line corresponds to $F_{\delta \delta}^{\prime} / H_{\delta \delta}$. The sum of the two ratios is equal to unity for all values of the measurement strength $\theta$.


FIG. 3. (Color online) The red region corresponds to values of the parameters $(\theta, \delta)$, where $F_{\phi \phi}^{\prime} / H_{\phi \phi}>1 / 2$ for $\phi=0$ (in turn, the white region corresponds to $F_{\delta \delta}^{\prime} / H_{\delta \delta}>1 / 2$ ).
and the final strong measurement by the same angle $\phi$, this result shows that the measurement scheme presented herein is optimal: one can always saturate the bound by using a two-step adaptive method, which allows us to measure the parameters near the optimal working point [12]. As we show in Fig. 2, in the case of $\phi=0$, one can tune the measurement strength $\theta$ in order to explore the estimation tradeoff and decide the amount of information desired for each of the two parameters. The value of the phase diffusion can only slightly change the behavior of the tradeoff between the two ratios corresponding to each parameter. In particular, while for small values of $\delta$ the range of values of $\theta$ where one can get more information about the phase $\phi$ is much larger than that corresponding to more information about the diffusion $\delta$, the tradeoff gets more balanced for $\delta \geqslant 1$. This behavior is displayed in Fig. 3, where we show the region of values $(\theta, \delta)$ for which the ratio $F_{\phi \phi}^{\prime} / H_{\phi \phi}$ is larger than $1 / 2$.

## III. EXPERIMENT

We illustrate the concepts developed in the previous section via a characterization of a weak measurement device based on the interferometric scheme introduced by Iinuma and colleagues [30]. We perform a weak measurement of the Stokes parameter $\mathcal{S}_{2}=|H\rangle\langle H|-|V\rangle\langle V|$ of a light beam, followed by a standard measurement of $\mathcal{S}_{2}=|H+V\rangle\langle H+V|-|H-V\rangle\langle H-V|$. We then tackle the problem of estimating the linear polarization direction, i.e., a phase shift on the basis of the circular polarizations, together with its associated noise.

A weak measurement necessitates an ancillary meter system to which the observable of interest is coupled. In the original experiment, the path inside a Mach-Zehnder interferometer acts as the ancillary system: the beam is first split on a polarizing beam splitter (PBS) and then recombined on a beam splitter after manipulation of the polarization [30,31]. In our arrangement, both the initial and the final beam


FIG. 4. (Color online) The measurement device. (a) Experimental setup: this is a modification over the original scheme of Iinuma et al. [30,31], in which $\mathcal{S}_{1}$ is coupled weakly to the path inside the Sagnac interferometer, thus to the two spatial exits. Detection is performed by means of calorimetric power meters; the probabilities $p(-,+)$ and $p(-,-)$ have not been resolved, their sum has been measured instead. (b) calibration curves for $p(+,+)+p(+,-)-$ $p(-,+)-p(-,-)$ associated with $\sigma_{z}$ (blue squares and solid line) and $[p(+,+)-p(+,-)] /[p(+,+)+p(+,-)]$ (red dots and solid line) associated with $\sigma_{x}$ as functions of the HWP1 angle $\phi$.
splitters of the interferometer are PBSs, and that makes it possible to adopt a stable Sagnac configuration instead of the Mach-Zehnder arrangement [Fig. 4(a)].

The input polarization of a cw $\mathrm{He}: \mathrm{Ne}$ laser is prepared by the first half-wave-plate (HWP1) as an arbitrary linear polarization $c_{H}|H\rangle+c_{V}|V\rangle$. The beam is then divided in two by a PBS, whose outputs constitute the two arms of the Sagnac interferometer. HWP2 is present in both modes, and it sets the measurement strength through its angle $\omega$ with respect the horizontal, while HWP3 is kept with its axis at $-45^{\circ}$ to ensure that the two spatial modes are coupled differently to the polarization. With this choice, the effect of the overall device on the polarization is

$$
\left(\begin{array}{cc}
\cos (2 \omega) & 0  \tag{11}\\
0 & \sin (2 \omega)
\end{array}\right)\binom{c_{H}}{c_{V}}
$$

for output 1 and

$$
\sigma_{x}\left(\begin{array}{cc}
\sin (2 \omega) & 0  \tag{12}\\
0 & \cos (2 \omega)
\end{array}\right)\binom{c_{H}}{c_{V}}
$$

for output 2. Therefore, we obtain an implementation of the measurement operators (6) up to a unitary; the orientation of the axis of HWP2 is linked to the setting of the measurement by $\theta=\frac{\pi}{2}-4 \omega$, which is set in our experiment to $\theta_{0}=58^{\circ}$, and it approaches the strong regime. Due to technical limitations, we only consider a three-outcome measurement without resolving $\sigma_{x}$ on one of the outputs. This results in a slight decrease of the available Fisher information, but, notably, the correlations between the two parameters, which are our main interest for this investigation, remain unaffected.

As a first step, we perform a calibration of our device by injecting pure states with known phases set by the angle of HWP1, and we register the outcome probability. We then obtain calibration curves, such as those in Fig. 4(b), using an interpolation [32]; arbitrary pairs of phase and dephasing can then be recovered by finding the values of $\phi$ and $\delta$ which are more compatible with these curves with a minimal residual approach. Following the technique in [33], in our experiment we have simulated the state (1) by injecting two pure states and adding the output signals with proper weighting. Figure 5


FIG. 5. (Color online) Experimental multiparameter estimation. The mixed state (1) is simulated by adding the data corresponding to the pure state with $\theta \simeq 3.18$ and $\theta \simeq 3.18-\pi$, with weights $\frac{1+e^{-\delta_{0}^{2}}}{2}$ and $\frac{1-e^{-\delta_{0}^{2}}}{2}$, respectively. For each point, the expected value of the dephasing is set to a value $\delta_{0}$, and the parameters are estimated by a minimal residual based on the curves in Fig. 2, implemented by a maximum likelihood algorithm; this delivers the values $\phi$ and $\delta$. The dashed curve shows the expected behavior $\delta=\delta_{0}$. The uncertainties are estimated with a Monte Carlo procedure: we simulate 10000 repetitions of the experiment in which the intensities vary within the experimental uncertainty.
reports the experimental values $\phi$ and $\delta$ obtained when varying the expected dephasing $\delta_{0}$. As expected, the estimated phase does not change with the amplitude of the dephasing, while we observe a small bias in the estimation of $\delta$, due to the fact that the data sets do not have the same intensity.

In a multiparameter estimation experiment, an important aspect is the manifestation of correlations between the estimated parameters. Since the main source of uncertainty in our experiment is the discretization imposed by the intensity meter, a repetition of the experiment would not lead to an insight into the errors on each parameter, nor into their correlations. To obtain a visualization of such uncertainties, we have employed a Monte Carlo routine that simulates repetitions of the estimation experiment by varying the intensity values within the range of the experimental uncertainties (see Fig. 6). This constitutes a simple and direct way to inspect the experimental covariance of the measured parameters. We then compare these histograms with the expected covariance $\Sigma=F / M^{\prime}$ for an ideal instrument. We emphasize that, due to the origin of the uncertainties in our measurement, the parameter $M^{\prime}$ should not be interpreted directly as an effective number of experiments, but rather as a scaling parameter to make the two figures comparable [34]. As expected, the shape of the correlations depends on the working value of $\delta_{0}$, and the trend is followed even in the presence of imperfections.

## IV. CONCLUSIONS

We have discussed the performance of weak measurements as a scheme for the joint estimation of a phase parameter and the amplitude of contextual dephasing. We have illustrated how these measurements obey the same bound as all POVMs, and we related the information obtained on the two parameters to the measurement setting. An experimental investigation has


FIG. 6. (Color online) Histograms of the Monte Carlo realizations for (a) $\delta_{0}=0.094$ and (b) $\delta_{0}=0.25$, compared with the expected covariances, as a function of the deviations $\tilde{\phi}$ and $\tilde{\delta}$ from the mean values. The covariance is calculated for $M^{\prime}=4 \times 10^{5}$; it serves as a comparison of the correlations between the two estimates, and it does not correspond to the Cramér-Rao bound, since $M^{\prime}$ is lower than the average photon number as an effect of extra noise. We observe that qualitatively the variation in the correlations follows the prediction for an ideal weak measurement device. It is worth remarking that when working at low $\delta_{0} \lesssim 0.05$, the optimization procedure that extracts $\{\phi, \delta\}$ sometimes gets trapped in a local minimum corresponding with $\delta=0$.
shown how these considerations remain valid qualitatively even in the presence of experimental imperfections. Our results highlight how these measurements represent a practical scheme for such a multiparameter scenario, but they present no intrinsic advantage.

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## APPENDIX: MULTIPARAMETER QUANTUM ESTIMATION THEORY

Let us consider a physical system parametrized by a set of $n$ parameters $\lambda=\left\{\lambda_{1}, \ldots, \lambda_{n}\right\}$. An estimator $\hat{\lambda}=\left\{\hat{\lambda}_{1}, \ldots, \hat{\lambda}_{n}\right\}$ is defined as a map from a set of $M$ measurement outcomes $X=\left\{x_{1}, \ldots, x_{M}\right\}$ to the possible values of the parameters $\lambda$. We consider unbiased estimators, i.e., estimators that, for a sufficiently large number of measurements $M$, satisfy the property $\mathbb{E}[\hat{\lambda}]=\lambda(\mathbb{E}[\cdot]$ denoting the average over all the possible measurement results $X$ ). The performances of unbiased estimators are quantified in terms of the covariance matrix $\Sigma$, whose elements are defined as

$$
\begin{equation*}
\Sigma_{i j}=\mathbb{E}\left[\left(\hat{\lambda}_{i}-\lambda_{i}\right)\left(\hat{\lambda}_{j}-\lambda_{j}\right)\right] \tag{A1}
\end{equation*}
$$

such that, for unbiased estimators, the diagonal elements correspond to the variances and off-diagonal elements to covariances. The covariance matrix is lower bounded according to the Cramér-Rao bound as

$$
\begin{equation*}
\Sigma \geqslant \frac{\mathbf{F}^{-1}}{M} \tag{A2}
\end{equation*}
$$

where $\mathbf{F}$ denotes the Fisher information matrix, with components

$$
\begin{equation*}
F_{i j}=\int d x p(x \mid \lambda)\left[\partial_{\lambda_{i}} \log p(x \mid \lambda)\right]\left[\partial_{\lambda_{j}} \log p(x \mid \lambda)\right] . \tag{A3}
\end{equation*}
$$

In quantum mechanics, the conditional probability entering in Eq. (A3) can be recast as $p(x \mid \lambda)=\operatorname{Tr}\left[\varrho_{\lambda} \Pi_{x}\right]$, where $\varrho_{\lambda}$ is the quantum state parametrized in terms of $\lambda$, and $\left\{\Pi_{x}\right\}$ is the POVM corresponding to the measurement performed. By defining the symmetric logarithmic derivative (SLD) operators by the equation $L_{k} \varrho_{\lambda}+\varrho_{\lambda} L_{k}=2 \partial_{\lambda_{k}} \varrho_{\lambda}$, one can introduce the quantum Fisher information (QFI) matrix $\mathbf{H}$, with elements

$$
\begin{equation*}
H_{i j}=\operatorname{Re}\left(\operatorname{Tr}\left[\varrho_{\lambda} L_{i} L_{j}\right]\right) . \tag{A4}
\end{equation*}
$$

The QFI matrix provides a more general bound for the covariance matrix, holding for any possible POVM, i.e.,

$$
\begin{equation*}
\Sigma \geqslant \frac{\mathbf{H}^{-1}}{M} \tag{A5}
\end{equation*}
$$

While the classical bound (A2) can always be saturated, for example with a maximum likelihood estimator, the quantum bound is proven to be always achievable only for a single parameter, with the optimal measurement corresponding to the eigenstates of the SLD operator. In the multiparameter case, the optimal measurements for each parameter (still corresponding to the eigenstates of the different SLD operators) may not
commute, and as a consequence the bound cannot in general be saturated. Moreover, different bounds can be obtained by
considering different derivation operators [35]; in our case, the SLD provides the most informative bound.
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