# Sex ratio and time to pregnancy: analysis of four large European population surveys 

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#### Abstract

Objective To test whether the secondary sex ratio (proportion of male births) is associated with time to pregnancy, a marker of fertility. Design Analysis of four large population surveys. Setting Denmark and the United Kingdom. Participants 49506 pregnancies. Main outcome measure Secondary sex ratio. Results No association was found between the sex ratio and time to pregnancy and no discernible trend was found for sex ratio with time to pregnancy, either within individual datasets or in the pooled analysis. The odds ratios were 0.97 ( $95 \%$ confidence interval 0.90 to 1.04) for contraceptive failures, 1.01 ( 0.96 to 1.05) for time to pregnancy of 2-4 months, 1.02 ( 0.97 to 1.08 ) for 5-10 months, 0.98 ( 0.93 to 1.03 ) for 11 months or more, and 0.88 ( 0.74 to 1.06 ) for fertility treatment, with 0-1 months as the reference category. Conclusion No association was found between the secondary sex ratio and time to pregnancy.


## INTRODUCTION

Evidence exists of a decline in semen quality over recent decades and a deterioration in the reproductive health of males, with an increase in the incidence of testicular cancer and probably of hypospadias and cryptorchid-ism-the "testicular dysgenesis syndrome". ${ }^{1}$ One cause that has been suggested is endocrine disruption due to environmental agents with oestrogenic or antiandrogenic effects, although direct evidence is lacking and other mechanisms are possible. ${ }^{1}$ Other things being equal, a decline in semen quality is expected to be followed by a declining trend in fertility, although during the same period no such decline-as measured by the time to pregnancy-has been observed. ${ }^{2}$

In parallel, during the same period the proportion of male offspring (commonly called the secondary sex ratio) has been declining in many countries, including the United States and Canada and several European countries. ${ }^{34}$ An association between fertility and secondary sex ratio could in principle result from universal biological processes (for example, the frequency and timing of intercourse relative to ovulation or levels of sex hormones in either partner) or from exposure to a toxin that affects both end points.

James has suggested that a low secondary sex ratio could be used as an indicator of risk factors to male reproductive health (as an alternative to semen quality, hormone levels or a functional measure of biological fertility such as time to pregnancy). ${ }^{5}$ Møller has suggested that an altered secondary sex ratio should be considered as part of the testicular dysgenesis syndrome, partly on the grounds of parallel trends but also because a case-control study of men with testicular cancer found an excess of female offspring (albeit of borderline significance) born two years or more before the father's diagnosis. ${ }^{4}$ This observation was confirmed in a study of Danish men born during 1945-80. ${ }^{6}$ The same association was not, however, observed at the aggregate level, in that Finland has a low incidence of testicular cancer compared with the other Nordic countries but also a low secondary sex ratio. ${ }^{4}$ Moreover, among 15218 men who attended the Sperm Analysis Laboratory in Copenhagen for various reasons (most commonly infertility), no association was found between the secondary sex ratio and any semen characteristics. ${ }^{7}$ Exposure to certain substances, such as the nematocide dibromochlorpropane ${ }^{8}$ and dioxin ${ }^{9}$ and related substances, ${ }^{1011}$ has been found to affect the secondary sex ratio and male fertility in a parallel way.

This parallelism of trend taken together with shared chemical determinants raises the question of whether the probability of conceiving a male child is linked to the parents' fertility. Such an association would suggest that these chemical agents, or others with a similar mode of action, may be causing the observed parallel trends in the secondary sex ratio and male fertility. In other words, the association would link the secondary sex ratio with the testicular dysgenesis syndrome and could lead towards elucidating the underlying cause of the secondary sex ratio and fertility, whether hormonal or not.

We tested the hypothesis that the secondary sex ratio is associated with the time to pregnancy, a marker of fertility, ${ }^{12}$ using four high quality datasets from previous population surveys.

## METHODS

We used datasets from four major European studies with population based samples: the Odense prenatal
study (Denmark), ${ }^{13}$ the Asclepios project (an occupational study), ${ }^{14}$ the Office for National Statistics omnibus study (representative of Great Britain), ${ }^{2}$ and the Millennium cohort study (representative of the United Kingdom). ${ }^{15}$
We identified 49506 liveborn singletons conceived without fertility treatment. Information on conception was obtained by questionnaire, which has been shown to be a valid, even after 20 years of recall. ${ }^{16}$ We classified pregnancies into contraceptive failures and those with a time to pregnancy of 0-1 months (reference category), 2-4 months, 5-10 months, and 11 months or more. Data on a further 496 babies born after fertility treatment were available and we included these as an additional category in the pooled analysis across all datasets (numbers were too small for inclusion of this category in the study specific analyses). Data on contraceptive failures were not reliably available from the Millennium cohort questionnaire.
We used logistic regression to examine the association between pregnancy categories and secondary sex ratio, adjusting for birth order, age of both parents (grouped), and study. To account for the lack of independence of observations due to about $40 \%$ of couples in the Odense dataset having more than one child we included a random effect for each couple in the model. To implement this we used penalised quasilikelihood ${ }^{17}$ through the command glmmPQL in the software SPLUS.

## RESULTS

Table 1 lists the details of the four major European surveys providing datasets. The figure shows the distribution of pregnancy categories for each study. Some differences between studies are apparent, which may reflect variations in couple fertility in time or space.

No association between secondary sex ratio and degree of fertility was observed for any of the studies separately or when pooled (table 2). For the pooled analysis the odds ratios were 0.97 ( $95 \%$ confidence interval 0.90 to 1.04 ) for contraceptive failures and 1.01 ( 0.96 to 1.05 ) for time to pregnancy of 2-4 months, 1.02 ( 0.97 to 1.08 ) for 5-10 months, and 0.98 ( 0.93 to 1.03 ) for 11 months or more, with $0-1$ months as the reference category. The odds of having a male offspring after successful fertility treatment were reduced ( $0.88,0.74$ to 1.06 ). Although the odds
ratios show small departures from 1.0 , the confidence intervals all include 1.0 and no trends can be seen for the individual studies or for the pooled analysis. The influence of clustering within couples was negligible, with almost identical results being produced for the analyses including random effects (table 2) and excluding random effects (results not shown).

An analysis using a dichotomous split of time to pregnancy corresponding to "clinical infertility" produced an odds ratio of 0.97 ( $95 \%$ confidence interval 0.93 to 1.03 ) for those with a time to pregnancy of 12 months or more compared with those classified as contraceptive failures or with a time to pregnancy of less than 12 months. A further analysis was carried out using time to pregnancy as a continuous variable, excluding contraceptive failures, fertility treated couples, and those with a time to pregnancy of more than 36 months, leaving 44209 babies. The exclusions prevented the small number of extreme values for time to pregnancy from influencing the fitted regression line. This produced a statistically non-significant regression coefficient ( $-0.0014,95 \%$ confidence interval -0.0042 to 0.0013 ).

## DISCUSSION

Using datasets from four large European studies we found no association between time to pregnancy (a marker of fertility) and secondary sex ratio (proportion of male births). No patterns were apparent to suggest a significant result if larger population numbers had been available.

Assuming no misclassification of time to pregnancy, we determined that the main analysis (excluding babies conceived after fertility treatment) had $99 \%$ power to detect a $4 \%$ difference in the percentage of male births between the top category (contraceptive failures) and bottom category (time to pregnancy 11 months or more). It also had a power of $93 \%$ to detect a difference of $3 \%$ between the categories, but a power of only $57 \%$ to detect a difference of $2 \%$. Although differences of $2 \%-4 \%$ in the secondary sex ratio represent large effects, these were none the less smaller than the $6.5 \%$ difference in proportion of male births found by one study, ${ }^{18}$ which compared couples with times to pregnancy of less than 12 months with those waiting at least 12 months.

Table 1|Characteristics of study populations

| Study (location) | Study design | Timing of data collection | Study period | Response rate (\%) | Respondents who had become parents (\%)* |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Odense prenatal study (Fyn, Denmark) | Pregnancy based | In antenatal clinic at around 20 weeks' gestation | 1972-97 | 92 | 100 |
| Asclepios project (European Union occupational study) | Cross sectional | Lifetime reproductive history; youngest child | 1995-8 | 57 | 86 |
| Office for National Statistics omnibus study (Great Britain) | Cross sectional | Lifetime reproductive history; first pregnancy | 1996 | 72 | 73 |
| Millennium cohort study (United Kingdom) | Pregnancy based | When child was aged about 9 months | 2001-2 | 82 | 52 |
| *Parents with a useable value for time to pregnancy. |  |  |  |  |  |

Table 2 | Male births in relation to contraceptive failures and time to pregnancy. Values are odds ratios ( $95 \%$ confidence intervals) unless stated otherwise

| Study | No of births | Contraceptive failures | Time to pregnancy (months)* |  |  |  | No of births after fertility treatment |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | 0-1 <br> (reference group) | 2-4 | 5-10 | $\geq 11$ |  |
| Odense prenatal study | 36674 | 0.98 (0.91 to 1.06) | 1.0 | 1.00 (0.95 to 1.06) | 1.00 (0.94 to 1.07) | 0.97 (0.91 to 1.04) | NA |
| Asclepios project | 1797 | 0.75 (0.52 to 1.08) | 1.0 | 0.93 (0.74 to 1.16) | 1.12 (0.82 to 1.52) | 1.03 (0.73 to 1.45) | 45 |
| Office for National Statistics omnibus study | 1817 | 0.92 (0.69 to 1.24) | 1.0 | 0.99 (0.77 to 1.27) | 1.18 (0.85 to 1.63) | 1.10 (0.82 to 1.48) | 60 |
| Millennium cohort study | 9218 | NA | 1.0 | 1.05 (0.95 to 1.16) | 1.08 (0.95 to 1.22) | 0.97 (0.86 to 1.11) | 391 |
| Pooled | 49506 | 0.97 (0.90 to 1.04) | 1.0 | 1.01 (0.96 to 1.05) | 1.02 (0.97 to 1.08) | 0.98 (0.93 to 1.03) | $\begin{gathered} 0.88 \text { (0.74 to } \\ 1.06) \dagger \\ \hline \end{gathered}$ |

Analyses are adjusted for birth order and parents' age groups, and allow for multiple pregnancies per couple. Pooled analysis is also adjusted for study.
NA=not available.
*Values greater than 1.0 indicate increased probability of male births compared with reference group. †Odds ratio ( $95 \%$ confidence interval) for 496 babies.

Even allowing for non-differential misclassification of time to pregnancy ${ }^{16}$ our study was of sufficient power to detect a difference in the proportion of male births similar to that found previously. ${ }^{18}$ If we assume a true underlying difference of $6.5 \%$ in the secondary sex ratio between couples with times to pregnancy of 12 months or less and a sensitivity of $79.9 \%$ and specificity of $94.9 \%$ for the binary indicator of time to pregnancy, ${ }^{16}$ then following the working in Armstrong et al ${ }^{19}$ we would still expect to observe a difference of over $4 \%$ in the proportion of male births between the two groups after allowing for misclassification.

Although several biases may occur in studies of times to pregnancy if they are not designed and analysed well, ${ }^{12}$ these primarily affect the design, analysis, and interpretation of studies using time to pregnancy as the outcome rather than as a predictor, as in our study. Furthermore, these biases are unlikely to be affected by the sex of the baby. We therefore believe that this is a truly negative study.

Few studies have been published on the possible association between fertility and the probability of a male offspring. One study found that in a population exposed to pesticides males tended to be conceived more quickly than females. ${ }^{20}$ Another study found a


Distribution of times to pregnancy by dataset
deficit of males among pregnancies that took more than a year to achieve. ${ }^{21}$ On the other hand, Smits et al found an increasing proportion of males born with longer waiting times to pregnancy. ${ }^{18}$ These contradictory findings are based on relatively small studies, and so the lack of an association we found using a larger dataset is not incompatible with this earlier literature; there could also be other negative studies that have not been published. Indeed, two large unpublished analyses have been posted on the BMJ website, ${ }^{2223}$ both of which found no association. Together with our study the sample size is almost 130000 . Although time to pregnancy is affected only by severely abnormal semen quality, this combined sample should have been sufficient to detect an association if one were present, even if it was smaller than that found by Smits et al. ${ }^{18}$ Nevertheless, the associations seen in the published studies could be real ${ }^{182021}$ : they could have resulted if the populations that they were based on had been exposed to an environmental agent that affected both the secondary sex ratio and time to pregnancy.

We were motivated to test our hypothesis because of parallel historical trends in the secondary sex ratio and health of the male reproductive system, as manifest by the testicular dysgenesis syndrome, which could indicate that whatever exposure is responsible for an increasing trend in this syndrome is also causing a reduction in the secondary sex ratio. This is supported by evidence from studies of exposed populations. Paternal occupational exposure to the nematocide dibromochlorpropane has been found to greatly impair male fertility, in some cases causing permanent sterility; the probability of a male birth was greatly reduced among those men who retained or recovered their fertility. ${ }^{8}$

In rodents even a low dose of dioxin impaired spermatogenesis. ${ }^{9}$ This outcome has been judged the critical end point of dioxin in that effects occur at a lower dose than for any other outcome. Certain polychlorinated biphenyls have dioxin-like effects, ${ }^{9}$ and more generally these and other agents that disrupt the

## WHAT IS ALREADY KNOWN ON THIS TOPIC

Semen quality and the probability of a male birth have both declined in recent decades The two may be linked, as certain chemical agents affect both of them

## WHAT THIS STUDY ADDS

No association was found between time to pregnancy and the secondary sex ratio (proportion of male offspring)
endocrine system may possibly cause impairment of the male reproductive system. ${ }^{11011}$ In humans, paternal exposure has been observed to affect the secondary sex ratio in some studies but not in others. ${ }^{1011}$
This perspective offers a reconciliation of our findings, those posted on the BMJ website, and those of Smits et al. Our study population covered several European countries, and the observed lack of association between time to pregnancy and the secondary sex ratio suggests that there is no relevant exposure common to all these countries. On the other hand, the population in Smits et al's study was drawn from a region in the Netherlands, which could be exposed to an environmental agent that has not so far been recognised. If so it would be a valuable clue in an area of research where such clues are urgently needed.
Alternatively, the secondary sex ratio and fertility could be linked through the female reproductive system, but the science here is unclear. Intercourse close to ovulation has been suggested to result in more males. ${ }^{24}$ As this would also lead to a relatively high probability of conception, male births would be associated with a shorter time before conception. Others, however, have suggested the opposite, on the basis of a different biological mechanism. ${ }^{25}$ We could find no evidence to support either of these hypotheses.

## CONCLUSION

Our findings of no evidence of an association between the secondary sex ratio and time to pregnancy concur with those of other large studies reported on the BMJ website. The pregnancies that we studied occurred during the past few decades in western Europe, where semen quality is thought to have declined ${ }^{1}$ in parallel with a decrease in proportion of male births. ${ }^{34}$ If these two phenomena were due to the same environmental agent (for example, dioxin or another agent that disrupts the endocrine system), we would expect to have found an association between relatively long times to pregnancy and a decrease in the secondary sex ratio. The lack of an association between the two in our study provides no evidence of a shared environmental cause. Neither do our findings offer any support for alternative hypotheses linking fertility and the secondary sex ratio through the female reproductive system.

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